

Matched Pair Tests with Nondetects

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Objectives for this webinar

1. Inform you of the challenges of computing statistical tests for paired data with nondetects
2. Inform you of what methods are available for paired data with nondetects.



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For More Information

on this topic and other methods for data analysis with nondetects, see our online course



Nondetects And Data Analysis
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Matched - Pairs Tests

<u>Group X</u>		<u>Group Y</u>
Observation 1	↔	Observation 1
Observation 2	↔	Observation 2
Observation 3	↔	Observation 3
Observation 4	↔	Observation 4

There is a direct relationship between
observations in the same row



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Pairing in Environmental Studies

Most commonly by time or location

		Urban	Ag
→ Jan			
→ Feb			
→ March			
→ April			

“Blocks”



Tests are run by computing the differences between the two values in the same row. No differences computed between values in different rows.

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Pairing in Environmental Studies

Most commonly by time or location

		Old Method	New Method
→ Well 1			
→ Well 2			
→ Well 3			
→ Well 4			

“Blocks”



Upgradient/Downgradient, Pre-event/Post-event are other common paired columns.

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Matched Pair Tests for censored data

Test	Distributional Assumption?	Tests for ____ of Paired Differences
Censored Paired	Parametric	mean
Sign Test	Nonparametric	median
Signed - Rank	Nonparametric	percentiles
PPW	Nonparametric	percentiles



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Example: Matched Pairs

- Mercury in soils was measured at the same sites in 1996 before a major fire, and after the fire in 2001 (Eppinger et al. , 2003).
- Measurements were 'blocked' by location. This minimizes causes of change other than the difference between the two years, which is attributed to the effect of the fire.

Q1: Is the mass of mercury over the site before the fire different than after the fire (a two-sided test on mean of paired differences)? **Parametric**

Q2: Are concentrations consistently higher or lower after the fire (a two-sided test on the percentiles of differences)? **Nonparametric**



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The Data

Hg in Soils before (X1996) and after (X2001) a major fire.

"Stacked" format

```
> head(SedHg)
  Mercury Year CenHg
    0.020 X1996     1
    0.021 X1996     0
    0.020 X1996     1
    0.020 X1996     1
    0.026 X1996     0
    0.020 X1996     1
    ...
```

"Unstacked" format

```
> head(EppsedHg)
  X2001 X1996 Cens01 Cens96
    0.02 0.020     1      1
    0.02 0.021     0      0
    0.05 0.020     0      1
    0.02 0.020     0      1
    0.02 0.026     1      0
    0.02 0.020     0      1
    ...
```



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Interval-censored differences (one of the challenges)

Pattern of censored concentrations (" $<$ " added to view) results in interval-censored paired differences (`mindiff`, `maxdiff`):

X2001	Cens01	X1996	Cens96	<code>mindiff</code>	<code>maxdiff</code>
<0.02	TRUE	<0.020	TRUE	-0.020	0.020
0.02	FALSE	0.021	FALSE	-0.001	-0.001
0.05	FALSE	<0.020	TRUE	0.030	0.050
0.02	FALSE	<0.020	TRUE	0.000	0.020
<0.02	TRUE	0.026	FALSE	-0.026	-0.006
0.10	FALSE	<0.020	TRUE	0.080	0.100



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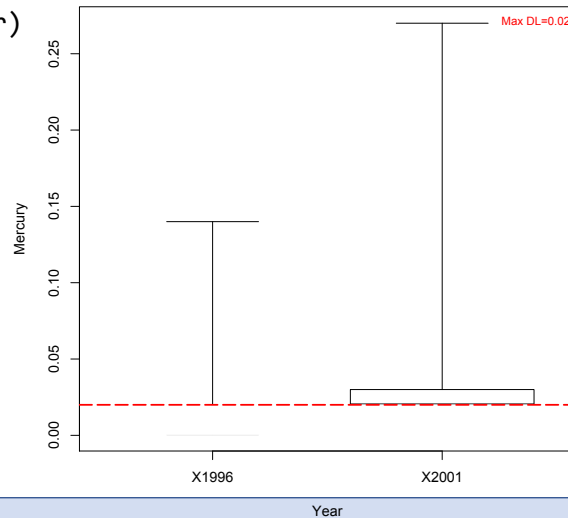
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Individual Boxplots of The Two Columns of Data

```
> cboxplot(Mercury, CenHg, Year)
```

Group boxplots DO NOT
show the effect of pairing
of samples at the same
site.

The 2001 data appear to
be higher than 1996 in
the upper percentiles



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The Sign Test

Null Hypothesis: Probability [2001 Hg > 1996 Hg] = one-half
(median difference = 0)

Alt. Hypothesis: Probability [2001 > 1996] does not \neq one-half
(two-sided)

Alt. Hypothesis: Probability [2001 > 1996] > one-half
(median difference > 0) (one-sided)

The sign test looks only at the +/-, not the magnitude, of the differences.



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Sign Test Computation

(censoring -> lots of ties)

Pattern of censored concentrations (" $<$ " added to view) results in several ties:

X	CensX	Y	CensY	Sign(X-Y)
<0.05	TRUE	<0.020	TRUE	0 (tie)
0.05	FALSE	0.021	FALSE	+
0.05	FALSE	<0.020	TRUE	+
0.03	FALSE	<0.050	TRUE	0 (tie)
<0.02	TRUE	0.026	FALSE	-
0.10	FALSE	<0.020	TRUE	+



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Two-sided Sign Test

using the `cen_signtest` script in `NADAscripts.R`

```
> cen_signtest(X2001, Cens01, X1996, Cens96)
Censored sign test for (x:X2001 - y:X1996) equals 0
  alternative hypothesis: true difference X2001 - X1996 not = 0
  n = 82   n+ = 55   n- = 8   ties: 19

  No correction for ties:  p-value = 9.76e-10 (many ties: is too small)
  Fong correction for ties: p-value = 0.0013
```



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The Signed-Rank Test

The test ranks the differences $x - y = d$ in absolute value, then places the algebraic sign \pm back onto the ranks.

Null Hypothesis: Median difference = 0; the two groups have similar cdfs

Alt. Hypothesis: Median difference is $\neq 0$ (two-sided); the cdfs are not the same

Alt. Hypothesis: Median difference is > 0 (one-sided); frequently $X > Y$

The signed-rank test generally has more power than the sign test because it uses the magnitudes of differences – a larger difference is stronger evidence than a smaller difference.

The test also requires a correction for ties -- these two tests assumed that ties were not present and their p-values are only correct if few ties are present. Censoring results in many ties.



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Censored Signed-Rank Test in R

Two-sided test:

```
> cen_signedrank.test (X2001, Cens01, X1996, Cens96)
```

```
Censored signed-rank test for x:X2001 - y:X1996 equals 0
```

```
alternative hypothesis: true difference X2001 - X1996 does not equal 0
```

```
Pratt correction for ties
```

```
n = 82    Z= 5.673    p-value = 0.00000001407
```

Other corrections have been proposed. Is this one the best one? It is the only one available in R (in the coin package). With a p-value this small, we reject the null hypothesis and state that mercury concentrations for the two years are not the same.



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Paired Prentice Wilcoxon Test

- A nonparametric test (O'Brien and Fleming, *Biometrics* 43, p. 169-180, 1987)
- Similar goal to the signed-rank test, to determine if one of the paired groups has "a shift in location" -- that y has consistently larger or smaller values than x.
- Has more power than the signed-rank test and sign tests when the differences are asymmetric (which environmental data often are).
- Has a good 'pedigree' of theory as part of a class of tests validated for censored data (linear rank tests on Prentice-Wilcoxon scores).
- Is built to handle ties. Doesn't need the retro-fit corrections of the sign and signed-rank tests for frequent ties.
- Can be thought of as something like a paired t-test on the scores.
- In terms of power and theoretical justification, I'd rank the usefulness of the three tests for testing censored environmental data as
PPW > signed-rank > sign test



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The PPW Test in R

Two-sided test:

```
> ppw.test(X2001, Cens01, X1996, Cens96)
      Paired Prentice-Wilcoxon test
```

data: X2001 and X1996

Paired Prentice Z = 6.044, n = 82, p-value = 1.504e-09

alternative hypothesis: true difference is not equal to 0

Median difference is 0.015

With a p-value this small, we reject the null hypothesis and state that mercury concentrations for the two years are not the same.



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Censored Paired Test

(a parametric test)

Null Hypothesis: mean difference = 0

Alt. Hypothesis: mean difference $\neq 0$ (two-sided)

Alt. Hypothesis: mean difference > 0 (one-sided)

If the mean difference is not 0, then the means of the two separate columns are not equal.



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Assumptions of the Censored Paired Test

- Primary assumption is that the paired differences have a normal distribution.
- Check this with a Q-Q plot for interval-censored data
- If non-normal, p-values may be too high (will not be too low)
- If p-values are low (0.01 etc.) then the violation of the normal assumption hasn't obscured the nonzero mean difference



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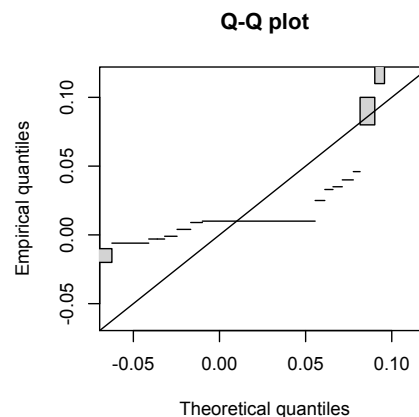
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Q-Q plot for interval-censored data

If the data do not follow a pattern similar to the straight line representing a normal distribution, the test's p-value may be too high.

Shaded rectangles show an area where the quantile (percentile) can be anywhere within the range of the box, due to the interval-censored nature for some differences.

If you need to transform the data, transform the X and Y original variables. Cannot transform the differences -- cannot take a log of a negative difference, for example.



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Censored Paired test in R: 2-sided

```
> cen_paired(X2001, Cens01, X1996, Cens96)
```

Censored paired test for mean(X2001 - X1996) equals 0.
alternative hypothesis: true mean difference does not equal 0.

n = 82 Z= 4.4747 p-value = 0.000007653

Mean difference = 0.02093835

the p-value is very low, so non-normality is not causing a problem here



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Comparing the mean of censored data to a standard

- Instead of a second column of data, insert a single number (the standard value) in the place of the y variable in cen_paired.
- To determine whether the mean exceeds a standard, set alternative = "greater".
- Is the same process used by computing a one-sided lower confidence interval on the mean in section 7b. If the LCL95 is above the standard then the mean is shown to exceed the standard with 95% confidence, and correspondingly the p-value for the test will be less than $(1-0.95) = 0.05$.



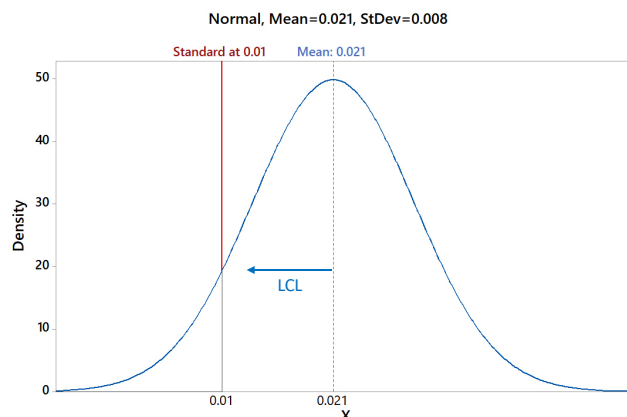
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The Censored Paired Test can be used to compare data to a standard

- This test is computed by MLE assuming a normal distribution.
- The methods in Section 7b allow you to test whether the mean exceeds a standard while assuming a lognormal or gamma distribution, which are often a better fit than is the normal distribution.



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Comparing the mean of censored data to a standard

Does the mean of 2001 Hg concentrations exceed a 'standard' of 0.01?

```
> cen_paired(X2001, Cens01, 0.01, alt = "greater")
Censored paired test for mean(X2001) equals 0.01
alternative hypothesis: true mean X2001 exceeds 0.01.
```

n = 82 Z= 5.2293 p-value = 8.508e-08

Mean X2001 = 0.03608

The p-value is small, so reject the null hypothesis of equality. The mean exceeds the standard.



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Summary: Tests for Difference Between Matched Pairs

Parametric and nonparametric tests answer two different questions.
Which did you want to answer?

Nonparametric: The Hg concentrations were more often higher in 2001 at a site than in 1996. (*Signed-rank, PPW or sign tests*)

Parametric: The mass of Hg over the entire site is higher in 2001 than in 1996. (*Censored Paired Test*)



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How Many Observations Are Censored Data Worth?

Your data set has 30 observations, 12 of which are nondetects. How many equivalent observations do you have for the information content possessed in these data? Is it more than just $30 - 12 = 18$ detected observations? How much more?

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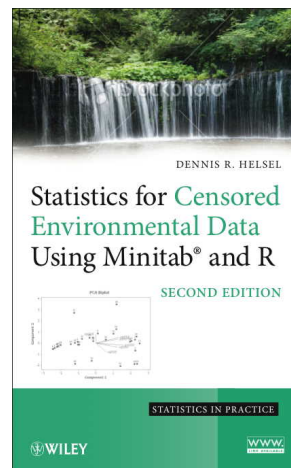
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For more on stats for data with NDs:

Statistics for Censored Environmental Data (the second edition)

by Dennis R. Helsel
(2012)



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- Some of the material is based on my book
[Statistics For Censored Environmental Data](#) by Dennis Helsel (2012).
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